# PROBABILISTIC GREEKS 

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#### Abstract

In the Black-Scholes-Merton option pricing formulas the coefficients that multiply the main variables (the price of the underlying and the strike price) are equal to some "Greeks" (partial derivatives of the price with respect to the main variables). In this paper we prove that this property is not only true for a log-normal distribution, but it is also satisfied by any distribution that comply with some natural conditions and by some exotic options. These identities are derived from a new integral representation of the Greeks. This representation allows to derive Greeks in an easy and systematic way simplifying the long computation of partial derivatives traditionally involved in obtaining them. When computing Greeks, these results can be applied to simplify the derivation of closed form expressions, to speed up numerical methods, and to obtain better accuracy.


## Resumen

En la fórmula de valuación de opciones Black-Scholes-Merton, los coeficientes que multiplican a las principales variables (el precio del subyacente y el precio de ejercicio) son iguales a algunas "Griegas" (derivadas parciales del precio con respecto a las principales variables). En este trabajo probamos que esta propiedad no es sólo verdadera para una distribución log-normal, sino también se satisface para cualquier distribución que cumpla con algunas condiciones naturales y para algunas opciones exóticas. Estas identidades son derivadas a partir de una nueva representación integral de las Griegas. Esta representación permite determinar las Griegas en una forma sencilla y sistemática simplificando las complejidades computacionales y matemáticas tradicionalmente relacionadas en estos calculos. Cuando calculamos las griegas, estos resultados pueden ser aplicados para simplificar la derivación de expresiones cerradas, acelerar los métodos numéricos y obtener mejores ajustes.

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## 1. Introduction

We will use the following standard notation: $S(t)$ or $S_{t}$ denote the price of a financial asset at time $t \in[0, T], q$ is the continuously compounded rate of dividends that a holder of this asset receives, $C(t)$ denotes the time $t$ value of a call option whose payoff function at expiration $T$ is $(S(T)-K)^{+}=\max \{S(T)-$ $K, 0\}$, where $K$ is the exercise price. Similarly, $P(t)$ denotes the time $t$ value of a put option whose payoff function at expiration $T$ is $(K-S(T))^{+}=\max \{K-$ $S(T), 0\}$. More generally, $V(t)$ denotes the time $t$ value of a closed portfolio (i.e., such that no value is added or taken from it). Also, $r$ is the continuously compounded risk free rate of interest.

The theory of option pricing (see Karatzas, I. and S. Shreve (1998) or Lamberton, D. and B. Lapeyere (1997)) establishes that in an arbitrage free market, under some standard assumptions:

$$
\begin{equation*}
V(t)=\mathrm{E}_{t}\left[V(T) e^{-r(T-t)}\right] \tag{1}
\end{equation*}
$$

for a given probability distribution and family of sigma algebras $F_{t}, 0 \leq t \leq T$. Where, $\mathrm{E}_{t}$ and $\mathrm{Pr}_{t}$ denote the expectation and probabilities taken with this distribution, and conditioning on $F_{t}$. We will use the usual name Risk Neutral Valuation for prices based on formula (1).

Many closed form formulas for European options can be presented as:

$$
\begin{align*}
& C(t)=c_{S} S(t)+c_{K} K  \tag{2}\\
& P(t)=p_{K} K+p_{S} S(t) \tag{3}
\end{align*}
$$

We will refer to $c_{S}, c_{K}, p_{S}$ and $p_{K}$ as the coefficients of the corresponding formula. This is a slight abuse of notation since they themselves depend on $S(t)$ and $K$. One single coefficient is undefined, but the pair $c_{S}, c_{K}$ is defined as two values that substituted in equation (2) give the value of the call (similarly for puts).

Let $A=\{w: S(T, w) \geq K\}$, let $A^{c}$ be its complement, and let $1_{A}$ and $1_{A^{c}}$ be their characteristic functions. Then, these coefficients admit the following representation:

$$
\begin{gather*}
c_{S}=\mathrm{E}_{t}\left[\frac{S(T) e^{-r(T-t)}}{S(t)} 1_{A}\right],  \tag{4}\\
c_{K}=-\mathrm{E}_{t}\left[e^{-r(T-t)} 1_{A}\right],  \tag{5}\\
p_{S}=-\mathrm{E}_{t}\left[\frac{S(T) e^{-r(T-t)}}{S(0)} 1_{A}\right],  \tag{6}\\
p_{K}=\mathrm{E}_{t}\left[e^{-r(T-t)} 1_{A}\right] \tag{7}
\end{gather*}
$$

If interest rates are assumed to be deterministic (a frequently used approximation for derivatives on stock and foreign exchange markets), formulas (5) and (7) can be expressed as $-e^{-r(T-t)} \operatorname{Pr}_{t}[A]$ and $e^{-r(T-t)} \operatorname{Pr}_{t}\left[A^{c}\right]$. We keep the
more general expressions because they are also valid when interest rates are stochastic.

These identities hold under any model where derivatives can be valued using the Risk Neutral Universe (i.e., where equation (1) is satisfied). Many variations of them are well known, some applications and generalizations can be found in Crouhy, M. et al. (2000), Geman, H. et al. (1995) and Gerber, H. and E. Shiu (1996). Since they are crucial for the results presented here, a proof is provided in appendix A.

Consider, now, the Black-Scholes-Merton model (see Black, F. and M. Scholes (1973) and Merton, R. (1973)). We will refer to it as B-S-M. Finding the expectations (4), (5), (6) and (7) under this model the B-S-M equations are obtained:

$$
\begin{gather*}
C(t)=S(t) e^{-q(T-t)} N\left(d_{1}\right)-K e^{-r(T-t)} N\left(d_{2}\right),  \tag{8}\\
P(t)=K e^{-r(T-t)} N\left(-d_{2}\right)-S(t) e^{-q(T-t)} N\left(-d_{1}\right),  \tag{9}\\
d_{1}=\frac{\ln \left(\frac{S(t)}{K}\right)+\left(r-q+\frac{\sigma^{2}}{2}\right)(T-t)}{\sigma \sqrt{(T-t)}}, \\
d_{2}=d_{1}-\sigma \sqrt{(T-t)} .
\end{gather*}
$$

Where $N($.$) is the cumulative distribution function of the standard normal$ distribution.

Consider now $\Delta_{\text {call }}(t)=\partial C(t) / \partial S(t)$. In the B-S-M model, after a laborious differentiation of formula (8) it turns out that $\Delta_{\text {call }}=e^{-q(T-t)} N\left(d_{1}\right)$, i.e. $\Delta_{\text {call }}$ equals the coefficient of the price. We prove here that the identity between $\Delta_{\text {call }}$ and the coefficient $c_{S}$ (or the corresponding expectation) holds under any distribution that satisfies a very natural condition. Similar results are shown for puts, some exotic options, and other Greeks.

The key property is that, while the random variable $\frac{S(T) e^{-r(T-i)}}{S(t)}$ depends on $\sigma$ and $T$, it is independent from the values of $S(t), K$ and $r$. So, the conditions required are:

$$
\begin{align*}
& \frac{S(T) e^{-r\left(T^{\prime}-t\right)}}{S_{t}} \text { is independent of } S_{t},  \tag{10}\\
& \frac{S(T) e^{-r(T-t)}}{S_{t}} \text { is independent of } K  \tag{11}\\
& \frac{S(T) e^{-r\left(T^{\prime}-t\right)}}{S_{t}} \text { is independent of } r \tag{12}
\end{align*}
$$

Condition (10) only says that we are interested in models of $S(T)$ that consider proportional growth rather than absolute growth, a necessary condition to model financial assets. Condition (11) says that the asset growth is independent of the strike price of an option written on it. This is satisfied in standard options where the strike price is fixed. Condition (12) says that, once
discounted, the asset price is independent of the interest rates. This is true in the Risk Neutral Universe (see Hull, J. (2000)).

These conditions are applied independently of each other to obtain equalities between Greeks and expectations. The fact that Greeks equal the corresponding coefficients has often been cited as a rule of thumb to remember the Greeks. This rule gets fully explained here. Its explanation is provided by lemma 1, below, and the fact that the first summand in the left hand side of the lemma vanishes for most financial derivatives.

The equality between coefficients and Greeks follows from a new integral representation of derivatives prices (proposition 1, below). Even though it is simple to obtain, it is very powerful. It allows to derive Greeks in an easy and systematic way simplifying the long computation of partial derivatives traditionally involved in obtaining them.

In section 2, the main results are stated. Section 3 provides some technical lemmas needed to obtain these results. In section 4, the results are proven.

## 2. Greeks as Expectations

In this section we prove the identities between $\Delta_{\text {call }}, \kappa_{\text {call }}$, and $\rho_{\text {call }}$ (derivatives of the call price with respect to $S_{t}, K$, and $r$ ) and the corresponding expectations. We need only to assume conditions (10), (11) and (12) respectively. We also state similar identities for $\Delta_{p u t}, \kappa_{p u t}$, and $\rho_{p u t}$. To prove them the same procedure, with minor modifications, can be followed. Alternatively, the results for puts can be deduced from the results for calls using the put-call parity.

Theorem 1

- If condition (10) is satisfied, then

$$
\begin{aligned}
\Delta_{c a l l} & =\mathrm{E}_{t}\left[\frac{S_{T} e^{-r(T-t)}}{S_{t}} 1_{A}\right] \\
\Delta_{p u t} & =-\mathrm{E}_{t}\left[\frac{S_{T} e^{-r(T-t)}}{S_{t}} 1_{A^{\mathrm{c}}}\right] .
\end{aligned}
$$

- If condition (11) is satisfied, then

$$
\begin{aligned}
\kappa_{\text {call }} & =-\mathrm{E}_{t}\left[e^{-r(T-t)} 1_{A}\right], \\
\kappa_{\text {put }} & =\mathrm{E}_{t}\left[e^{-r(T-t)} 1_{A^{c}}\right]
\end{aligned}
$$

- If condition (12) is satisfied, then

$$
\begin{aligned}
\rho_{\text {call }} & =(T-t) K \mathrm{E}_{t}\left[e^{-r(T-t)} 1_{A}\right], \\
\rho_{\text {put }} & =-(T-t) K \mathrm{E}_{t}\left[e^{-r(T-t)} 1_{A^{c}}\right] .
\end{aligned}
$$

Using the representation of coefficients as expectations (equations (4), (5), (6) and (7)), this theorem can also be expressed in terms of the coefficients. Obtaining, thus,

Corollary 1

- If conditions (10) and (11) are satisfied, then (2) is satisfied with $c_{S}=\Delta_{\text {call }}$ and $c_{K}=\kappa_{\text {call }}$ and (3) is satisfied with $p_{S}=\Delta_{p u t}$ and $p_{K}=\kappa_{\text {put }}$.
- If conditions (10) and (12) are satisfied, then (2) is satisfied with $c_{S}=$ $\Delta_{\text {call }}$ and $c_{K}=-\rho_{\text {call }} K^{-1}(T-t)^{-1}(3)$ is satisfied with $p_{S}=\Delta_{p u t}$ and $p_{K}=-\rho_{p u t} K^{-1}(T-t)^{-1}$.
There is a simple and general probabilistic expression for $\Gamma_{\text {call }}=\partial^{2} C(t) / \partial S(t)^{2}$ and for $\Gamma_{p u t}=\partial^{2} P(t) / \partial S(t)^{2}$ :

Theorem 2

- If condition (10) is satisfied, then

$$
\begin{aligned}
\Gamma_{c a l l} & =\frac{K e^{-r\left(T^{\prime}-t\right)}}{S_{t}} f_{S_{T}}(K), \\
\Gamma_{p u t} & =\frac{K e^{-r(T-t)}}{S_{t}} f_{S_{T}}(K) .
\end{aligned}
$$

Where, as usual, $f_{X}(x)$ denotes the density function of the random variable $X$ applied to $x$. Note that under the B-S-M model, using the corresponding log-normal distribution for $f_{S_{T}}$, and after some algebraic manipulations, the usual expression for $\Gamma$ is obtained (see Wilmott, P (2000)). These results will be derived from the following representation of the price of a derivative. Its proof is provided in appendix $B$.
Proposition 1
Let

$$
y_{0}=\frac{K e^{-r(T-t)}}{S_{t}} \text { and } Y=\frac{S(T) e^{-r(T-t)}}{S_{t}}
$$

then

$$
\begin{align*}
& C(t)=\int_{y_{0}}^{\infty}\left(S_{t} y-K e^{-r(T-t)}\right) \mathrm{d} F_{Y}(y),  \tag{13}\\
& P(t)=\int_{-\infty}^{y_{0}}\left(S_{t} y-K e^{-r(T-t)}\right) \mathrm{d} F_{Y}(y) . \tag{14}
\end{align*}
$$

Where $F_{X}(x)$ denotes the cumulative distribution function of the random variable $X$, applied to $x$, and conditioning on $F_{t}$ (i.e. assuming that $S_{t}$ is known).

## 3. Calculus Lemmas

To compute $\Delta_{\text {call }}$, we need to differentiate equation (13). Permuting the derivative and the integral is not possible because $y_{0}$ depends on $S_{0}, K$, and $r$. In this section we develop results that take care of this obstacle. For this, we will use the following Lemma (see Haaser, N. et al. (1964)).

Lemma 1
Let $f(x, y)$ be a continuous function on $\left[\alpha_{1}, \alpha_{2}\right] \times\left[\beta_{1}, \beta_{2}\right]$ such that $\frac{\partial f(x, y)}{\partial y}$ exists and is continuous. Let $a$ and $b$ be two differentiable functions on
$\left[\beta_{1}, \beta_{2}\right]$ such that for every $y \in\left[\beta_{1}, \beta_{2}\right], a(y) \in\left[\alpha_{1}, \alpha_{2}\right]$ and $b(y) \in\left[\alpha_{1}, \alpha_{2}\right]$. Then,

$$
\begin{equation*}
\frac{\partial\left(\int_{a(y)}^{b(y)} f(x, y)\right)}{\partial y} \mathrm{~d} x=\int_{a(y)}^{b(y)} \frac{\partial(f(x, y))}{\partial y} \mathrm{~d} x+f(b(y), y) b^{\prime}(y)-f(a(y), y) a^{\prime}(y) \tag{15}
\end{equation*}
$$

In the applications, after a change of variables, $b(y)$ will be the strike price of a financial derivative and $f$ the payoff of a derivative. Also, only one of the integration limits will depend on $y$. In this case the second and third term of the right hand side of equation (15) vanish. For ease of reference we state this simplified result.
Lemma 2
Assume that $a, b$ and $f$ satisfy the conditions of Lemma 1. If $a$ is constant and $f(b(y), y)=0$, then

$$
\begin{equation*}
\frac{\partial \int_{a}^{b(y)} f(x, y)}{\partial y} \mathrm{~d} x=\int_{a}^{b(y)} \frac{\partial f(x, y)}{\partial y} \mathrm{~d} x . \tag{16}
\end{equation*}
$$

If $b$ is constant and $f(a(y), y)=0$, then

$$
\begin{equation*}
\frac{\partial \int_{a(y)}^{b} f(x, y)}{\partial y} \mathrm{~d} x=\int_{a(y)}^{b} \frac{\partial f(x, y)}{\partial y} \mathrm{~d} x . \tag{17}
\end{equation*}
$$

## 4. Proofs of the Theorems

Proof of Theorem 1. We proceed now to compute $\Delta_{\text {call }}$ differentiating the integral of (13), applying lemma 2 , and using condition (10):

$$
\begin{align*}
\Delta_{c a l l} & =\frac{\partial C(t)}{\partial S_{t}} \\
& =\frac{\partial}{\partial S_{t}}\left(\int_{y_{0}}^{\infty}\left(S_{t} y-K e^{-r(T-t)}\right) \mathrm{d} F_{Y}(y)\right) \\
& =\int_{y_{0}}^{\infty} \frac{\partial\left(S_{t} y-K e^{-r(T-t)}\right)}{\partial S_{t}} \mathrm{~d} F_{Y}(y) \\
& =\int_{y_{0}}^{\infty} y \mathrm{~d} F_{Y}(y)  \tag{18}\\
& =\int_{K}^{\infty}\left(\frac{S_{T} e^{-r(T-t)}}{S_{t}}\right) \mathrm{d} F_{S_{T}}\left(S_{T}\right) \\
& =\mathrm{E}_{t}\left(\frac{S_{T} e^{-r(T-t)}}{S_{t}}-1_{A}\right)
\end{align*}
$$

Similarly, we can obtain $\kappa_{\text {call }}$ assuming (11).

$$
\begin{aligned}
\kappa_{\text {call }} & =\frac{\partial C(t)}{\partial K} \\
& =\frac{\partial}{\partial K}\left(\int_{y_{0}}^{\infty}\left(S_{t} y-K e^{-r(T-t)}\right) \mathrm{d} F_{Y}(y)\right) \\
& =\int_{y_{0}}^{\infty} \frac{\partial\left(S_{t} y-K e^{-r(T-t)}\right)}{\partial K} \mathrm{~d} F_{Y}(y) \\
& =\int_{y_{0}}^{\infty}-e^{-r(T-t)} \mathrm{d} F_{Y}(y) \\
& =\int_{K}^{\infty}-e^{-r(T-t)} \mathrm{d} F_{S_{T}}\left(S_{T}\right) \\
& =-\mathrm{E}_{t}\left(e^{-r(T-t)} 1_{A}\right)
\end{aligned}
$$

Assuming (12), we can also obtain $\rho_{\text {call }}$.

$$
\begin{aligned}
\rho_{\text {call }} & =\frac{\partial C(t)}{\partial r} \\
& =\frac{\partial}{\partial r}\left(\int_{y_{0}}^{\infty}\left(S_{t} y-K e^{-r(T-t)}\right) \mathrm{d} F_{Y}(y)\right. \\
& =\int_{y_{0}}^{\infty} \frac{\partial\left(S_{t} y-K e^{-r(T-t)}\right)}{\partial r} \mathrm{~d} F_{Y}(y) \\
& =\int_{y_{0}}^{\infty}-(T-t) K e^{-r(T-t)} \mathrm{d} F_{Y}(y) \\
& =-(T-t) K \mathrm{E}_{t}\left(e^{-r(T-t)} 1_{A}\right)
\end{aligned}
$$

Proof of Theorem 2. During the proof of theorem 1, in equation (18), we obtained the following equality $\Delta_{\text {call }}=\int_{y_{0}}^{\infty} y \mathrm{~d} F_{Y}(y)$. So ${ }^{1}$

$$
\begin{aligned}
\Gamma_{c a l l} & =\frac{\partial \Delta_{\text {call }}(t)}{\partial S_{t}} \\
& =\frac{\partial\left(\int_{y_{0}}^{\infty} y \mathrm{~d} F_{Y}(y)\right)}{\partial S_{t}} \\
& =\frac{\partial\left(\int_{y_{0}}^{\infty} y f_{Y}(y) \mathrm{d} y\right)}{\partial S_{t}}
\end{aligned}
$$

Since we are assuming condition (10), $Y$ is independent of $S_{t}$. The integral of the last expression only depends on $S_{t}$ through $y_{0}$, the limit of integration. So, in this case, we can just use the Fundamental Theorem of Calculus to obtain:

[^1]\[

$$
\begin{aligned}
\Gamma_{\text {call }} & =y_{0} f_{Y}\left(y_{0}\right) \frac{\mathrm{d} y_{0}}{\mathrm{~d} S_{t}} \\
& =\frac{K e^{-r(T-t)}}{S_{t}} f_{S_{T}}(K)
\end{aligned}
$$
\]

The last equality follows from the definition of $y_{0}$ and from the formula of change of variables for density functions.

## 5. Concluding Remarks

In this paper, we have proved that the partial derivatives of the option premium with respect to the price of the underlying and the strike price, main variables of Black-Scholes-Merton option pricing formulas, do not need the assumption of log-normal distribution to compute them, but they are also satisfied by any distribution that comply with some natural conditions. Then, we suggest a new way to derive Greeks from a new integral representation. This proposal follows two theorems, proven in the section 4 , that allow us to reduce the computational time to obtain them. Finally, we show that our results fullfill the features of the original Greeks.

## Appendices

## A. Coefficients as Expectations

Proof of equalities (4), (5), (6), and (7). To apply (1), we need a closed portfolios, so if $S(t)$ is the price of an asset that pays dividends, let $\widehat{S}(t)=S(t) e^{q t}$ be the portfolio resulting from reinvesting all dividends in the asset. Defining also $\widehat{K}=K e^{q t}$, the payoff function of the call is $(S(T)-K)^{+}=e^{-q T}(\widehat{S}(T)-\widehat{K})^{+}$. So the call on $S$ can be treated as a call on $\widehat{S}$, with strike price $\widehat{K}$, multiplied by a constant. This is the usual method to treat dividends and all the extensions that follow. (See Hull, J. (2000) and Merton, R. (1973)), Then

$$
\begin{aligned}
C(t) & =\mathrm{E}_{t}\left(e^{-q T}(\widehat{S}(T)-\widehat{K})^{+} e^{-r(T-t)}\right) \\
& =\mathrm{E}_{t}\left((S(T)-K)^{+} e^{-r(T-t)}\right) \\
& =\mathrm{E}_{t}\left((S(T)-K) e^{-r(T-t)} 1_{A}\right) \\
& =S(t) \mathrm{E}_{t}\left(\frac{S(T) e^{-r(T-t)}}{S(t)} 1_{A}\right)-K \mathrm{E}_{t}\left(e^{-r(T-t)} 1_{A}\right)
\end{aligned}
$$

The results follow from the last expression.

## B. The Integral Representation

Proof of proposition (1). Starting from equation (1), we have

$$
\begin{aligned}
C(t) & =\mathrm{E}\left((S(T)-K)^{+} e^{-r(T-t)}\right) \\
& =\int_{0}^{\infty}\left(S_{T}-K\right)^{+} e^{-r(T-t)} \mathrm{d} F_{S_{T}}\left(S_{T}\right) \\
& =\int_{K}^{\infty}\left(S_{T}-K\right) e^{-r(T-t)} \mathrm{d} F_{S_{T}}\left(S_{T}\right) \\
& =\int_{K}^{\infty}\left(S_{t}\left(\frac{S_{T} e^{-r(T-t)}}{S_{t}}\right)-K e^{-r(T-t)}\right) \mathrm{d} F_{S_{T}}\left(S_{T}\right) \\
& =\int_{y_{0}}^{\infty}\left(S_{t} y-K e^{-r(T-t)}\right) \mathrm{d} F_{Y}(y)
\end{aligned}
$$

The last equality follows from a simple change of variables. The proof for the put is similar.

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[^1]:    1 In the proof, we are assuming that $S_{T}$ follows a continuous distribution. The proof can also be done for a discrete distribution and is almost identical.

